

Markov Chains: An Introduction/Review

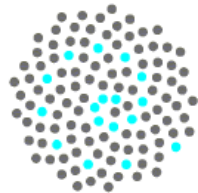
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THE UNIVERSITY
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AUSTRALIAN RESEARCH COUNCIL

Centre of Excellence for Mathematics
and Statistics of Complex Systems

Modified and shortened by Longin Jan Latecki
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Andrei A. Markov (1856 – 1922)



Random Processes

A random process is a collection of random variables indexed by some set I , taking values in some set S .

- I is the index set, usually time, e.g. \mathbb{Z}^+ , \mathbb{R} , \mathbb{R}^+ .
- S is the state space, e.g. \mathbb{Z}^+ , \mathbb{R}^n , $\{1, 2, \dots, n\}$, $\{a, b, c\}$.

We classify random processes according to both the index set (discrete or continuous) and the state space (finite, countable or uncountable/continuous).

Markov Processes

- A random process is called a *Markov Process* if, conditional on the current state of the process, its future is independent of its past.
- More formally, $X(t)$ is Markovian if has the following property:

$$\begin{aligned} & \mathbb{P}(X(t_n) = j_n \mid X(t_{n-1}) = j_{n-1}, \dots, X(t_1) = j_1) \\ &= \mathbb{P}(X(t_n) = j_n \mid X(t_{n-1}) = j_{n-1}) \end{aligned}$$

for all finite sequences of times $t_1 < \dots < t_n \in I$ and of states $j_1, \dots, j_n \in S$.

Time Homogeneity

A Markov chain $(X(t))$ is said to be *time-homogeneous* if

$$\mathbb{P}(X(s+t) = j \mid X(s) = i)$$

is independent of s . When this holds, putting $s = 0$ gives

$$\mathbb{P}(X(s+t) = j \mid X(s) = i) = \mathbb{P}(X(t) = j \mid X(0) = i).$$

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Probabilities depend on elapsed time, not absolute time.

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- Example: a frog hopping on 3 rocks. Put $S = \{1, 2, 3\}$.

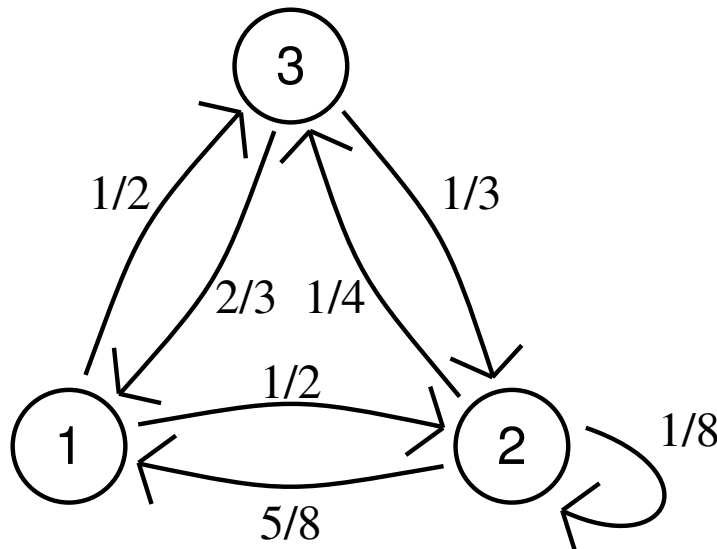
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- We can gain some insight by drawing a picture:



DTMCs: n -step probabilities

- We have P , which tells us what happens over one time step; let's work out what happens over two time steps:

$$\begin{aligned} p_{ij}^{(2)} &= \mathbb{P}(X_2 = j \mid X_0 = i) \\ &= \sum_{k \in S} \mathbb{P}(X_1 = k \mid X_0 = i) \mathbb{P}(X_2 = j \mid X_1 = k, X_0 = i) \\ &= \sum_{k \in S} p_{ik} p_{kj}. \end{aligned}$$

$$= \sum_k \mathbb{P}(X_2 = j, X_1 = k \mid X_0 = i)$$

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- So $P^{(2)} = PP = P^2$.
- Similarly, $P^{(3)} = P^2P = P^3$ and $P^{(n)} = P^n$.

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$$\pi_j^{(n)} = \sum_{k \in S} \mathbb{P}(X_0 = k) \mathbb{P}(X_n = j \mid X_0 = k)$$

$$\mathbb{P}\{X_n = j\}$$

//

$$= \sum_{k \in S} \pi_k^{(0)} p_{kj}^{(n)}$$

$$\sum_{k \in S} \mathbb{P}(X_0 = k, X_n = j)$$

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- Or, in matrix notation, $\pi^{(n)} = \pi^{(0)} P^n$; similarly we can show that $\pi^{(n+1)} = \pi^{(n)} P$.

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- We say that a state i leads to j (written $i \rightarrow j$) if it is possible to get from i to j in some finite number of jumps: $p_{ij}^{(n)} > 0$ for some $n \geq 0$.

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- The relation \leftrightarrow partitions the state space into *communicating classes*.
- We call the state space **irreducible** if it consists of a single communicating class.
- These properties are easy to determine from a transition probability graph.

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- We call a state i *recurrent* or *transient* according as $\mathbb{P}(X_n = i \text{ for infinitely many } n)$ is equal to one or zero.

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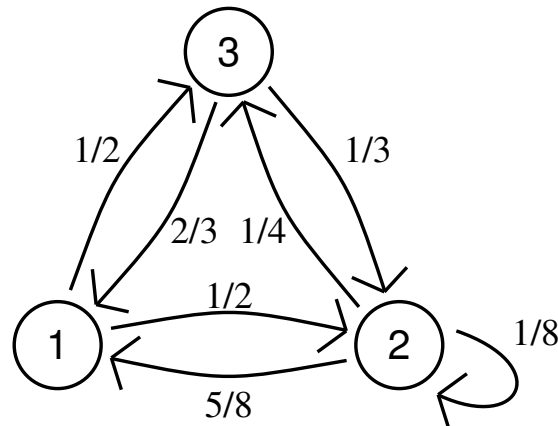
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- We also assume throughout that no states are *periodic*.

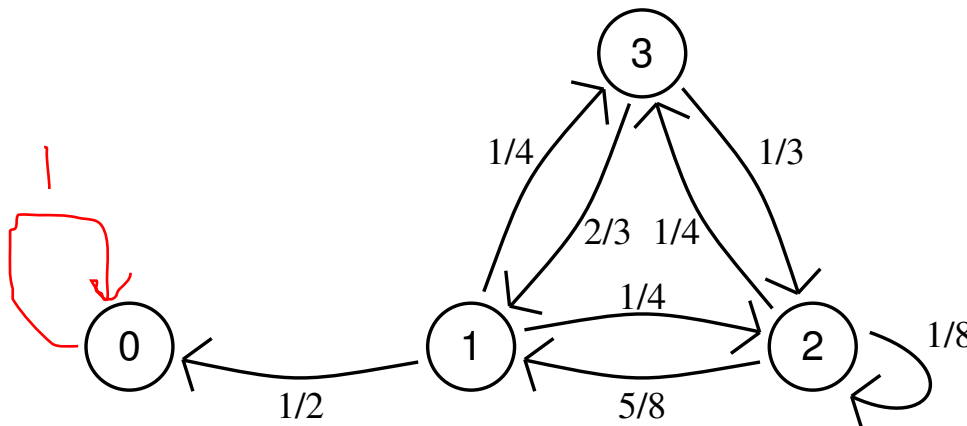
DTMCs: Two examples

- S irreducible:



$$P = \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{5}{8} & \frac{1}{8} & \frac{1}{4} \\ \frac{2}{3} & \frac{1}{3} & 0 \end{pmatrix}$$

- $S = \{0\} \cup C$, where C is a transient class:



$$P = \begin{pmatrix} 1 & 0 & 0 & 0 \\ \frac{1}{2} & 0 & \frac{1}{4} & \frac{1}{4} \\ 0 & \frac{5}{8} & \frac{1}{8} & \frac{1}{4} \\ 0 & \frac{2}{3} & \frac{1}{3} & 0 \end{pmatrix}$$

DTMCs: Quantities of interest

Quantities of interest include:

- Hitting probabilities.
- Expected hitting times.
- Limiting (stationary) distributions.
- Limiting conditional (quasistationary) distributions.

DTMCs: Hitting probabilities

Let α_i be the probability of hitting state 1 starting in state i .

• Clearly $\alpha_1 = 1$; and for $i \neq 1$,

$$\begin{aligned}\alpha_i &= \mathbb{P}(\text{hit } 1 \mid \text{start in } i) \\ &= \sum_{k \in S} \mathbb{P}(X_1 = k \mid X_0 = i) \mathbb{P}(\text{hit } 1 \mid \text{start in } k) \\ &= \sum_{k \in S} p_{ik} \alpha_k\end{aligned}$$

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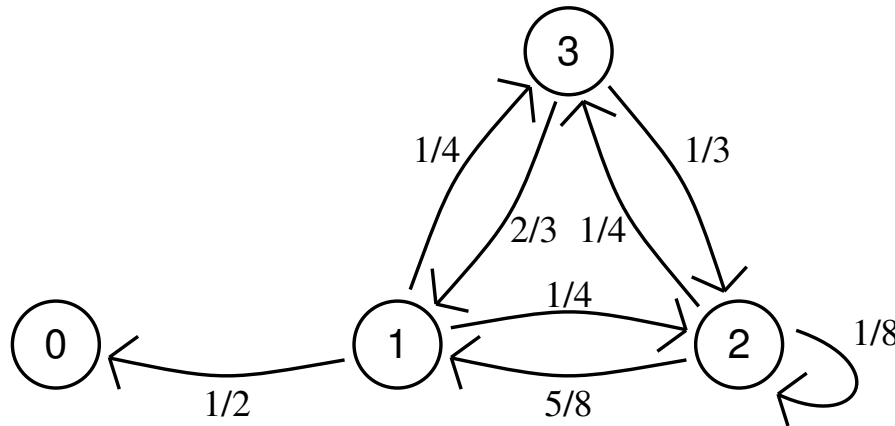
$X_1 = k$

$\implies \alpha = P\alpha, \alpha = (\alpha_1, \dots, \alpha_n)^T$

- Sometimes there may be more than one solution $\alpha = (\alpha_i, i \in S)$ to this system of equations.

If this is the case, then the hitting probabilities are given by the *minimal* such solution.

Example: Hitting Probabilities



$$P = \begin{pmatrix} 1 & 0 & 0 & 0 \\ \frac{1}{2} & 0 & \frac{1}{4} & \frac{1}{4} \\ 0 & \frac{5}{8} & \frac{1}{8} & \frac{1}{4} \\ 0 & \frac{2}{3} & \frac{1}{3} & 0 \end{pmatrix}$$

Let α_i be the probability of hitting state 3 starting in state i .

So $\alpha_3 = 1$ and $\alpha_i = \sum_k p_{ik} \alpha_k$:

$$\alpha_0 = \alpha_0$$

$$\alpha_1 = \frac{1}{2}\alpha_0 + \frac{1}{4}\alpha_2 + \frac{1}{4}\alpha_3$$

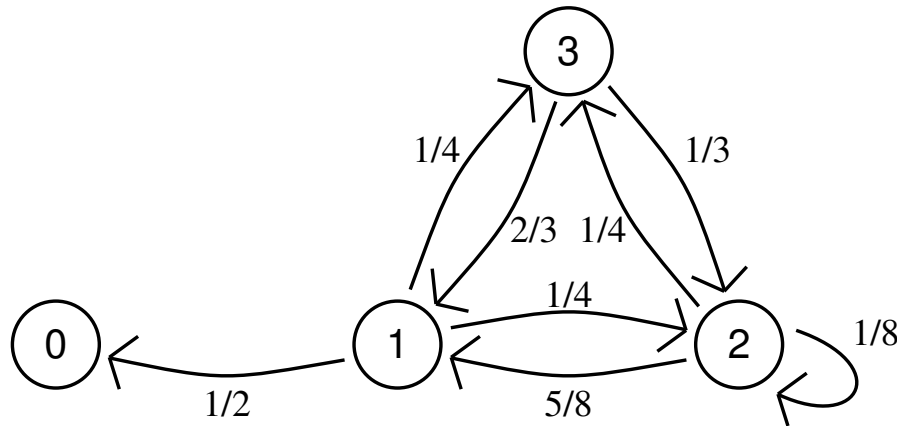
$$\alpha_2 = \frac{5}{8}\alpha_1 + \frac{1}{8}\alpha_2 + \frac{1}{4}\alpha_3$$

$$\alpha = \begin{pmatrix} \alpha_0 \\ \alpha_1 \\ \alpha_2 \\ \alpha_3 \end{pmatrix}$$

$$\alpha = P \alpha$$

relation to Label Propagation

Example: Hitting Probabilities



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Let α_i be the probability of hitting state 3 starting in state i .

$$\alpha = \begin{pmatrix} 0 \\ \frac{9}{23} \\ \frac{13}{23} \\ 1 \end{pmatrix} \approx \begin{pmatrix} 0 \\ 0.39 \\ 0.57 \\ 1 \end{pmatrix}.$$

DTMCs: The Limiting Distribution

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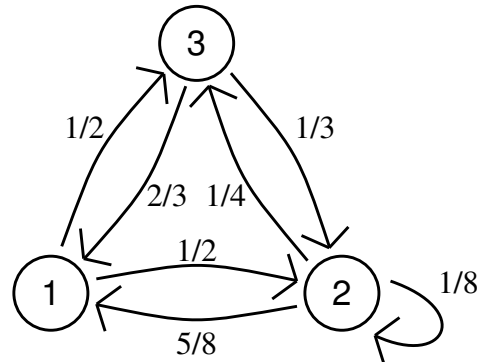
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(Such a distribution is called *stationary*.)

- This limiting distribution does not depend on the initial distribution.
- When the state space is infinite, it may happen that $\pi_j^{(n)} \rightarrow 0$ for all j .

Example: The Limiting Distribution



$$P = \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{5}{8} & \frac{1}{8} & \frac{1}{4} \\ \frac{2}{3} & \frac{1}{3} & 0 \end{pmatrix}$$

Substituting P into $\pi = \pi P$ gives

relation to PageRank

$$\pi_1 = \frac{5}{8}\pi_2 + \frac{2}{3}\pi_3,$$

$$\pi_2 = \frac{1}{2}\pi_1 + \frac{1}{8}\pi_2 + \frac{1}{3}\pi_3,$$

$$\pi_3 = \frac{1}{2}\pi_1 + \frac{1}{4}\pi_2,$$

which together with $\sum_i \pi_i = 1$ yields

$$\pi = \left(\frac{38}{97} \quad \frac{32}{97} \quad \frac{27}{97} \right) \approx \left(0.39 \quad 0.33 \quad 0.28 \right).$$

DTMCs: The Limiting Conditional Dist'n

Assume that the state space consists of an absorbing state and a transient class ($S = \{0\} \cup C$).

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- The limiting distribution is $(1, 0, 0, \dots)$.
- Instead of looking at the limiting behaviour of

$$\mathbb{P}(X_n = j \mid X_0 = i) = p_{ij}^{(n)},$$

we need to look at

$$\mathbb{P}(X_n = j \mid X_n \neq 0, X_0 = i) = \frac{p_{ij}^{(n)}}{1 - p_{i0}^{(n)}}$$

for $i, j \in C$.

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- It turns out we need a solution $m = (m_i, i \in C)$ of

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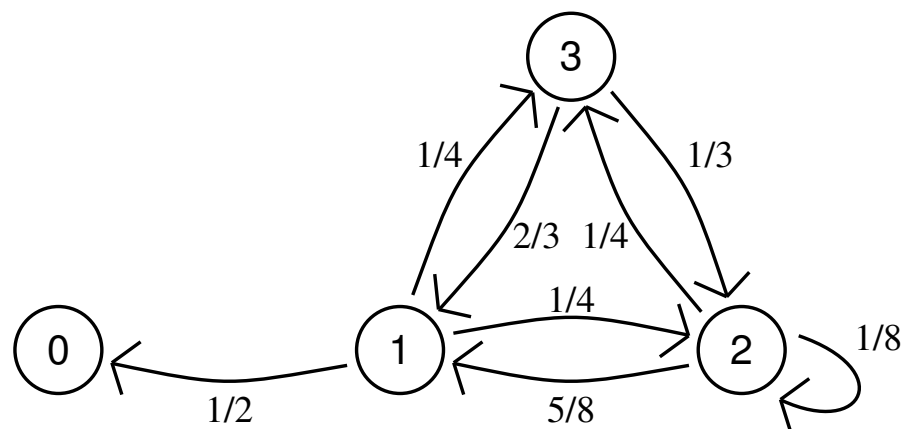
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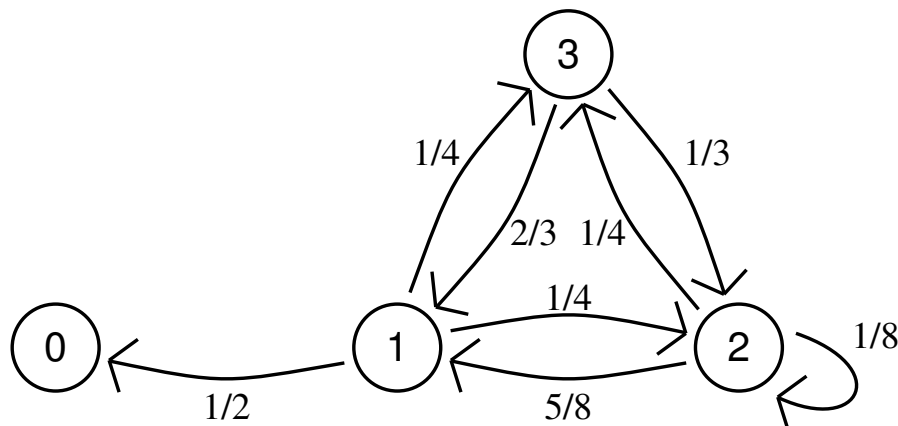
- If C is a finite set, there is a unique such r .
- If C is infinite, there is $r^* \in (0, 1)$ such that all r in the interval $[r^*, 1)$ are admissible; and the solution corresponding to $r = r^*$ is the LCD.

Example: Limiting Conditional Dist'n



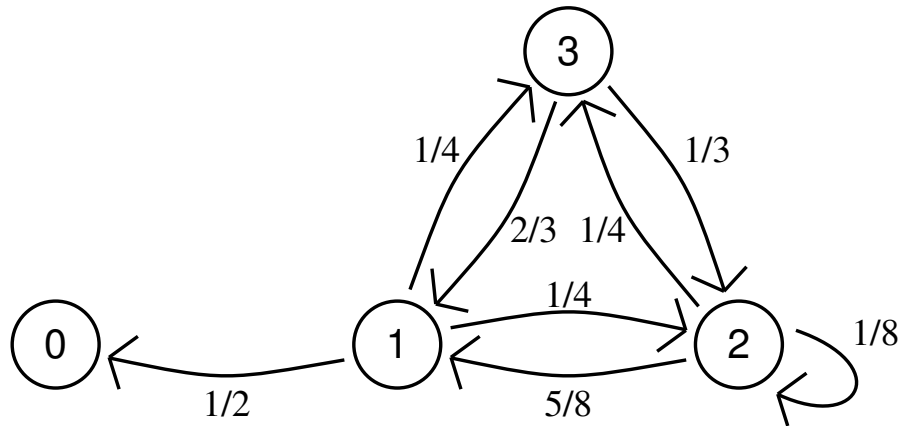
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$$P_C = \begin{pmatrix} 0 & \frac{1}{4} & \frac{1}{4} \\ \frac{5}{8} & \frac{1}{8} & \frac{1}{4} \\ \frac{2}{3} & \frac{1}{3} & 0 \end{pmatrix}$$

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Solving $mP_C = rm$, we get

$$r_1 \approx 0.773 \quad \text{and} \quad m \approx (0.45, 0.30, 0.24)$$

DTMCs: Summary

From the one-step transition probabilities we can calculate:

- n -step transition probabilities, ✓
- hitting probabilities, ✓
- expected hitting times, ✗
- limiting distributions, and ✓
- limiting conditional distributions. ✗